

# A NOTE ON ESSENTIAL SMOOTHNESS IN THE HESTON MODEL

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ABSTRACT. This note studies an issue relating to essential smoothness that can arise when the theory of large deviations is applied to a certain option pricing formula in the Heston model. The note identifies a gap, based on this issue, in the proof of Corollary 2.4 in [2] and describes how to circumvent it. This completes the proof of Corollary 2.4 in [2] and hence of the main result in [2], which describes the limiting behaviour of the implied volatility smile in the Heston model far from maturity.

## 1. INTRODUCTION

In [2] the authors study the limiting behaviour of the implied volatility in the Heston model as maturity tends to infinity. The main aim of this note is to give a rigorous account of the relationship between the concept of essential smoothness and the large deviation principle for the family of random variables  $(X_t/t \pm E_\lambda/t)_{t \geq 1}$ , where the process  $X$  denotes the log-spot in Heston model (5) and  $E_\lambda$  is an exponential random variable with parameter  $\lambda > 0$  independent of  $X$ . This note fills a gap in the proof of Corollary 2.4 in [2] and hence completes the proof of the main result in [2], which describes the limiting behaviour of the implied volatility smile in the Heston model far from maturity.

The note is organized as follows. Section 2 describes the relevant concepts of the large deviation theory and discusses how the effective domain changes when a family of random variables is perturbed by an independent exponential random variable. Section 3 discusses the failure of essential smoothness when the Heston model is perturbed by an independent exponential, which is what causes the gap in the proof of Corollary 2.4 in [2]. Section 3 also proves Theorem 3, which fills the gap.

## 2. THE LARGE DEVIATION PRINCIPLE FOR RANDOM VARIABLES IN $\mathbb{R}$

We briefly recall the basic facts of the large deviation theory in  $\mathbb{R}$  (see monograph [1, Ch. 2] for more details). Let  $(Z_t)_{t \geq 1}$  be a family of random variables with  $Z_t \in \mathbb{R}$ .  $J$  is a *rate function* if it is lower semicontinuous and  $J(\mathbb{R}) \subset [0, \infty]$  holds. The family  $(Z_t)_{t \geq 1}$  satisfies the *large deviation principle (LDP)* with the *rate function*  $J$  if for every Borel set  $B \subset \mathbb{R}$  we have

$$(1) \quad - \inf_{x \in B^\circ} J(x) \leq \liminf_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{P}[Z_t \in B] \leq \limsup_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{P}[Z_t \in B] \leq - \inf_{x \in \bar{B}} J(x),$$

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with the convention  $\inf \emptyset = \infty$  the relative notions of interior (interior  $B^\circ$ , closure  $\overline{B}$  and boundary  $\overline{B} \setminus B^\circ$  are in the topology of  $\mathbb{R}$ ).

The Gärtner-Ellis theorem (Theorem 1 below) gives sufficient conditions for a family  $(Z_t)_{t \geq 1}$  to satisfy the LDP (see monograph [1, Section 2.3] for details). Let  $\Lambda_t(u) := \log \mathbb{E} [e^{uZ_t}] \in (-\infty, \infty]$  be the cumulant generating function of  $Z_t$ . Assume that for every  $u \in \mathbb{R}$

$$(2) \quad \Lambda(u) := \lim_{t \rightarrow \infty} \Lambda_t(tu)/t \quad \text{exists in } [-\infty, \infty] \quad \text{and} \quad 0 \in \mathcal{D}_\Lambda^\circ,$$

where  $\mathcal{D}_\Lambda := \{u \in \mathbb{R} : \Lambda(u) < \infty\}$  is the *effective domain* of  $\Lambda$  and  $\mathcal{D}_\Lambda^\circ$  is its interior. The *Fenchel-Legendre transform*  $\Lambda^*$  of the convex function  $\Lambda$  is defined by the formula

$$(3) \quad \Lambda^*(x) := \sup\{ux - \Lambda(u) : u \in \mathbb{R}\} \quad \text{for } x \in \mathbb{R}.$$

Under the assumption in (2),  $\Lambda^*$  is lower semicontinuous with compact level sets  $\{x : \Lambda^*(x) \leq \alpha\}$  (see [1, Lemma 2.3.9(a)]) and  $\Lambda^*(\mathbb{R}) \subset [0, \infty]$  and hence satisfies the definition of a *good rate function*. We now state the Gärtner-Ellis theorem (see [1, Section 2.3] for its proof).

**Theorem 1.** *Let the random variables  $(Z_t)_{t \geq 1}$  satisfy the assumption in (2). If  $\Lambda$  is essentially smooth and lower semicontinuous, then LDP holds for  $(Z_t)_{t \geq 1}$  with the good rate function  $\Lambda^*$ .*

The function  $\Lambda : \mathbb{R} \rightarrow (-\infty, \infty]$  defined in (2) is *essentially smooth* if it is (a) differentiable in  $\mathcal{D}_\Lambda^\circ$  and (b) *steep*, i.e.  $\lim_{n \rightarrow \infty} |\Lambda'(u_n)| = \infty$  for every sequence  $(u_n)_{n \in \mathbb{N}}$  in  $\mathcal{D}_\Lambda^\circ$  that converges to a boundary point of  $\mathcal{D}_\Lambda^\circ$ . If  $\mathcal{D}_\Lambda^\circ$  is a strict subset of  $\mathbb{R}$ , which is the case in the setting of [2] (see also Section 3 below), essential smoothness, which plays a key role in the proof of Theorem 1, is not automatic.

The following question is of central importance in [2]: does the LDP persist if a family of random variables  $(Z_t)_{t \geq 1}$  is perturbed by an independent exponential random variable  $E_1$ ? It is implicitly assumed in the proof of Corollary 2.4 in [2] (see the last line on page 17 and lines 4 and 14 on page 18) that if  $(Z_t)_{t \geq 1}$  satisfies the assumptions of Theorem 1, then so do the families  $(Y_t^{1+})_{t \geq 1}$  and  $(Y_t^{1-})_{t \geq 1}$ , where  $Y_t^{1\pm} = Z_t \pm E_1/t$ , and the LDP is applied. In particular the authors in [2] assume that the limiting cumulant generating functions of  $(Y_t^{1\pm})_{t \geq 1}$  are essentially smooth. However the following simple lemma holds.

**Lemma 2.** *Let  $(Z_t)_{t \geq 1}$  satisfy the assumption in (2) with a limiting cumulant generating function  $\Lambda$ . Let  $\lambda > 0$  and  $E_\lambda$  an exponential random variable independent of  $(Z_t)_{t \geq 1}$  with  $\mathbb{E}[E_\lambda] = 1/\lambda$  and let  $Y_t^{\lambda\pm} := Z_t \pm E_\lambda/t$ . Then the families of random variables  $(Y_t^{\lambda\pm})_{t \geq 1}$  satisfy the assumption in (2) and the corresponding limiting cumulant generating functions are given by*

$$\Lambda^{\lambda+}(u) = \begin{cases} \Lambda(u), & \text{if } u \in \mathcal{D}_\Lambda \cap (-\infty, \lambda), \\ \infty, & \text{otherwise,} \end{cases} \quad \text{and} \quad \Lambda^{\lambda-}(u) = \begin{cases} \Lambda(u), & \text{if } u \in \mathcal{D}_\Lambda \cap (-\lambda, \infty), \\ \infty, & \text{otherwise.} \end{cases}$$

*Remarks. (a)* Let  $(Z_t)_{t \geq 1}$  satisfy the assumption in (2) and assume further that  $\Lambda$  is differentiable in  $\mathcal{D}_\Lambda^\circ$ . If  $1 \in \mathcal{D}_\Lambda^\circ$ , then the right-hand boundary point of the interior of the effective domain  $\mathcal{D}_{\Lambda^{1+}}^\circ$  is equal to 1 and Lemma 2 implies that the limiting cumulant generating function  $\Lambda^{1+}$  of  $(Y_t^{1+})_{t \geq 1}$  is

- neither essentially smooth, since  $\Lambda^{1+}$  is not steep at 1,
- nor lower semicontinuous at 1, since it is differentiable in  $\mathcal{D}_{\Lambda^{1+}}^\circ$  with  $\Lambda^{1+}(1) = \infty$ .

Loss of steepness and lower semicontinuity occurs also for  $(Y_t^{1-})_{t \geq 1}$  in the case where  $-1 \in \mathcal{D}_\Lambda^\circ$ .

**(b)** Lemma 2 implies that if  $(Z_t)_{t \geq 1}$  satisfies the assumptions of Theorem 1 and  $\mathcal{D}_\Lambda$  is contained in  $(-\infty, \lambda)$ , for some  $\lambda > 0$ , then  $(Y_t^{\lambda+})_{t \geq 1}$  also satisfies the assumptions of Theorem 1 and hence the LDP with a good rate function  $\Lambda^*$ . An analogous statement holds for  $(Y_t^{\lambda-})_{t \geq 1}$ .

*Proof.* Note that  $\log \mathbb{E} [e^{uE_\lambda}]$  is finite and equal to  $\log(\lambda/(\lambda - u))$  if and only if  $u \in (-\infty, \lambda)$ . For all large  $t$  and  $u \in \mathcal{D}_\Lambda \cap (-\infty, \lambda)$ , the assumption in (2) implies that  $\Lambda_t^{\lambda+}(tu) = \log \mathbb{E} [\exp(tuY_t^{\lambda+})]$  is finite and that the formula holds

$$(4) \quad \Lambda_t^{\lambda+}(tu) = \Lambda_t(tu) + \log \frac{\lambda}{\lambda - u}, \quad \text{where} \quad \Lambda_t(tu) = \log \mathbb{E} [\exp(tuZ_t)].$$

The inequality  $u \geq \lambda$  implies that, since  $\Lambda_t(tu) > -\infty$ , we have  $\Lambda_t^{\lambda+}(tu) = \infty$  for all  $t$  and hence  $\Lambda^{\lambda+}(u) = \infty$ . If  $u \in (\mathbb{R} \setminus \mathcal{D}_\Lambda) \cap (-\infty, \lambda)$ , then (4) yields  $\Lambda^{\lambda+}(u) = \lim_{t \nearrow \infty} \Lambda_t^{\lambda+}(tu)/t = \infty$ . This proves the lemma for  $(Y_t^{\lambda+})_{t \geq 1}$ . The case of  $(Y_t^{\lambda-})_{t \geq 1}$  is analogous.  $\square$

### 3. ESSENTIAL SMOOTHNESS CAN FAIL

The Heston model  $S = e^X$  is a stochastic volatility model with the log-stock process  $X$  given by

$$(5) \quad dX_t = -\frac{Y_t}{2} dt + \sqrt{Y_t} dW_t^1 \quad \text{and} \quad dY_t = \kappa(\theta - Y_t) dt + \sigma \sqrt{Y_t} dW_t^2,$$

where  $\kappa, \theta, \sigma > 0$ ,  $Y_0 = y_0 > 0$ ,  $X_0 = x_0 \in \mathbb{R}$  and  $W^1, W^2$  are standard Brownian motions with correlation  $\rho \in (-1, 1)$ . The standing assumption

$$(6) \quad \rho\sigma - \kappa < 0,$$

is made in [2] (see equation (2.2) in Theorem 2.1 on page 5 of [2]). In particular the inequality in (6) implies that  $S$  is a strictly positive true martingale and allows the definition of the share measure  $\tilde{\mathbb{P}}$  via the Radon-Nikodym derivative  $d\tilde{\mathbb{P}}/d\mathbb{P} = e^{X_t - x_0}$ .

The authors' aim in [2] is to obtain the limiting implied volatility smile as maturity tends to infinity at the strike  $K = S_0 e^{xt}$  for any  $x \in \mathbb{R}$  in the Heston model. Their main formula is given in Corollary 3.1 of [2]. A key step in the proof of [2, Corollary 3.1] is given by [2, Corollary 2.4]. In the proof of [2, Corollary 2.4] (see last line on page 17 and lines 4 and 14 on page 18) it is implicitly assumed that the LDP for  $(X_t/t)_{t \geq 1}$  implies the LDP for the family  $(X_t/t \pm E_1/t)_{t \geq 1}$ . However, as we have seen in Section 2 (see remarks following Lemma 2), Theorem 1 cannot be applied directly

to the family  $(X_t/t \pm E_1/t)_{t \geq 1}$ , even if  $(X_t/t)_{t \geq 1}$  satisfies its assumptions. We start with a precise description of the problem and present the solution in Theorem 3.

*Remarks.* (i) Under (6), a simple calculation implies that  $\Lambda$  and  $\mathcal{D}_\Lambda$  of the family  $(X_t/t)_{t \geq 1}$  are:

$$(7) \quad \Lambda(u) = -\frac{\theta\kappa}{\sigma^2} \left( u\rho\sigma - \kappa + \sqrt{\Delta(u)} \right) \quad \text{for } u \in \mathcal{D}_\Lambda \quad \text{and} \quad \mathcal{D}_\Lambda = [u_-, u_+] \quad \text{where}$$

$$(8) \quad u_\pm = \left( 1/2 - \rho\kappa/\sigma \pm \sqrt{(\kappa/\sigma - \rho)\kappa/\sigma + 1/4} \right) / (1 - \rho^2).$$

In (7) the function  $\Delta$  is a quadratic  $\Delta(u) = (u\rho\sigma - \kappa)^2 - \sigma^2(u^2 - u)$  and the boundary points  $u_+$  and  $u_-$  of the effective domain  $\mathcal{D}_\Lambda$  are its zeros. Elementary calculations show that  $\Lambda$  is essentially smooth and that the unique minimum of  $\Lambda^*$  is attained at  $\Lambda'(0) = -\theta/2$ . Therefore  $(X_t/t)_{t \geq 1}$  satisfies the LDP with the good rate function  $\Lambda^*$ , defined in (3), by Theorem 1.

(ii) Under the share measure  $\tilde{\mathbb{P}}$ , given by  $d\tilde{\mathbb{P}}/d\mathbb{P} = e^{X_t - x_0}$ , we have  $\tilde{\mathbb{E}}[e^{uX_t}] = e^{-x_0} \mathbb{E}[e^{(u+1)X_t}]$  for all  $u \in \mathbb{R}$  and  $t > 0$  and hence the family  $(X_t/t)_{t \geq 1}$  under  $\tilde{\mathbb{P}}$  satisfies the assumption in (2) with the limiting cumulant generating function  $\tilde{\Lambda}(u) = \Lambda(u+1)$ ,  $\mathcal{D}_{\tilde{\Lambda}} = [u_- - 1, u_+ - 1]$ . As before,  $(X_t/t)_{t \geq 1}$  satisfies the LDP under  $\tilde{\mathbb{P}}$  with the strictly convex good rate function  $\tilde{\Lambda}$ , which satisfies  $\tilde{\Lambda}^*(x) = \Lambda^*(x) - x$  for all  $x \in \mathbb{R}$  and attains its unique minimum at  $\tilde{\Lambda}'(0) = \Lambda'(1) = \theta\kappa/(\kappa - \rho\sigma)$ .

**Theorem 3.** *Let the process  $X$  be given by (5) and assume that (6) holds. Let  $E_1$  be the exponential random variable with  $\mathbb{E}[E_1] = 1$ , which is independent of  $X$ . Then the following limits hold:*

$$(9) \quad \lim_{t \nearrow \infty} \frac{1}{t} \log \mathbb{P} [X_t - x_0 + E_1 < xt] = -\Lambda^*(x) \quad \text{for } x \leq \Lambda'(0) = -\theta/2;$$

$$(10) \quad \lim_{t \nearrow \infty} \frac{1}{t} \log \tilde{\mathbb{P}} [X_t - x_0 - E_1 > xt] = x - \Lambda^*(x) \quad \text{for } x \geq \Lambda'(1) = \theta\kappa/(\kappa - \rho\sigma);$$

$$(11) \quad \lim_{t \nearrow \infty} \frac{1}{t} \log \tilde{\mathbb{P}} [X_t - x_0 - E_1 \leq xt] = x - \Lambda^*(x) \quad \text{for } x \in [\Lambda'(0), \Lambda'(1)];$$

where  $\Lambda$  is given in (7), its Fenchel-Legendre transform  $\Lambda^*$  is defined in (3) and  $d\tilde{\mathbb{P}}/d\mathbb{P} = e^{X_t - x_0}$ .

*Remark.* The limits in Theorem 3 are precisely the limits that arise in the proof of [2, Corollary 2.4] (see the last line on page 17 and lines 4 and 14 on page 18) and are claimed to hold since the family  $(X_t/t)_{t \geq 1}$  satisfies the LDP under  $\mathbb{P}$  and  $\tilde{\mathbb{P}}$  by Remarks (i) and (ii) above and Theorem 1. However Lemma 2 implies that the limiting cumulant generating function  $\Lambda^{1+}$  of the family of random variables  $(Z_t + E_1/t)_{t \geq 1}$ , where  $Z_t = (X_t - x_0)/t$ , is neither lower semicontinuous nor essentially smooth. Hence Theorem 1 cannot be applied to  $(Z_t + E_1/t)_{t \geq 1}$ . An analogous issue arises under the measure  $\tilde{\mathbb{P}}$ .

*Proof.* The basic idea of the proof is simple: for (9) we sandwich the probability  $\mathbb{P} [X_t - x_0 + E_1 < xt]$  between two tail probabilities of two families of random variables, which satisfy the LDP with the same rate function  $\Lambda^*$  by Lemma 2 and Theorem 1. The limits in (10) and (11) follow similarly.

For given parameter values in the Heston model pick  $\lambda > u_+$ , where  $u_+$  is defined in (8). Let  $E_\lambda$  be an exponential random variable with  $\mathbb{E}[E_\lambda] = 1/\lambda$ , defined on the same probability space as  $X$  and  $E_1$  and independent of both. Since  $u_+ > 1$ , we have the elementary inequality

$$(12) \quad \mathbb{P}[E_\lambda < \alpha] = I_{\{\alpha > 0\}} \left(1 - e^{-\lambda\alpha}\right) \leq I_{\{\alpha > 0\}} (1 - e^{-\alpha}) = \mathbb{P}[E_1 < \alpha] \quad \text{for any } \alpha \in \mathbb{R}.$$

The inequality

$$(13) \quad \mathbb{P}[X_t - x_0 + E_\lambda < xt] \leq \mathbb{P}[X_t - x_0 + E_1 < xt]$$

follows by conditioning on  $X_t$  and applying (12). On the other hand, since  $E_1 > 0$  a.s., we have

$$(14) \quad \mathbb{P}[X_t - x_0 + E_1 < xt] \leq \mathbb{P}[X_t - x_0 < xt].$$

Lemma 2 implies that the families of random variables  $(Z_t + E_\lambda/t)_{t \geq 1}$  and  $(Z_t)_{t \geq 1}$ , where  $Z_t = (X_t - x_0)/t$ , both have the limiting cumulant generating function equal to  $\Lambda$  given in (7) with the effective domain  $\mathcal{D}_\Lambda = [u_-, u_+]$ . Since  $\Lambda$  is essentially smooth and lower semicontinuous on  $\mathcal{D}_\Lambda$  and the assumption in (2) is satisfied, Theorem 1 implies that  $(Z_t + E_\lambda/t)_{t \geq 1}$  and  $(Z_t)_{t \geq 1}$ , satisfy the LDP with the good rate function  $\Lambda^*$ . Since  $x$  in (9) is assumed to be less or equal to the unique minimum  $\Lambda'(0) = -\theta/2$  of  $\Lambda^*$  (see Remark (i) above) and  $\Lambda^*$  is non-negative and strictly convex, the LDP (see the inequalities in (1)) and the inequalities in (13) and (14) imply the limit in (9).

To prove (10) pick  $\lambda > 1 - u_-$  and note that the inequality in (12) and conditioning on  $X_t$  yield

$$(15) \quad \tilde{\mathbb{P}}[X_t - x_0 > xt] \geq \tilde{\mathbb{P}}[X_t - x_0 - E_1 > xt] \geq \tilde{\mathbb{P}}[X_t - x_0 - E_\lambda > xt].$$

As before, Lemma 2 and Theorem 1 imply that  $(Z_t - E_\lambda/t)_{t \geq 1}$  and  $(Z_t)_{t \geq 1}$  satisfy the LDP with the convex rate function  $\tilde{\Lambda}^*$ , which by Remark (ii) above attains its unique minimum at  $\tilde{\Lambda}'(1) = \theta\kappa/(\kappa - \rho\sigma)$ . Since  $x \geq \tilde{\Lambda}'(1)$  in (10), the limit follows. A similar argument implies the limit in (11) for all  $x \in [\tilde{\Lambda}'(0), \tilde{\Lambda}'(1)]$ , which concludes the proof.  $\square$

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