

# Measures of Systemic Risk: Analysing CoVaR

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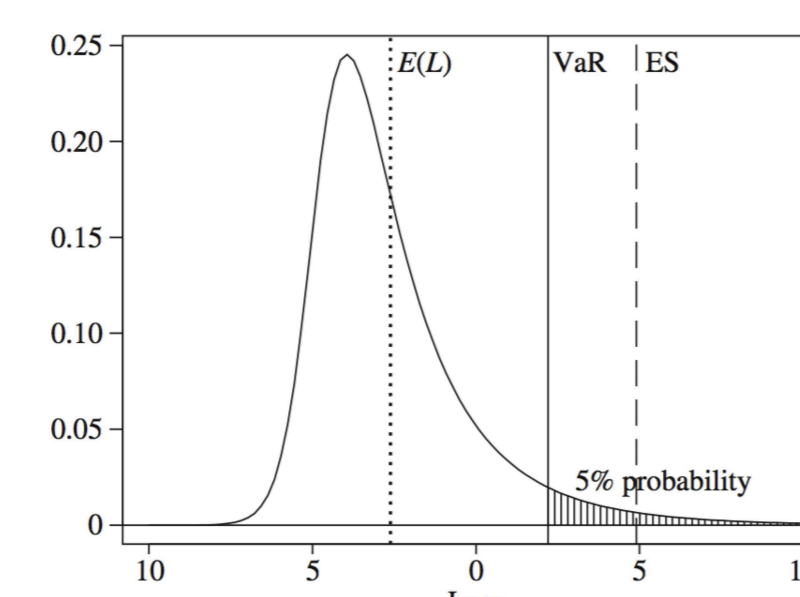
## Systemic Risk and Bank Capital Regulation

- Motivation: Account for the dependence structure between banks in the computation of their respective regulatory capitals;
- In 2008, Adrian and Brunnermeier introduced CoVaR – Conditional-Value-at-Risk – as a dependence adjusted version of VaR. Girardi and Ergün introduced a modified definition of CoVaR in an M-GARCH setting;
- Goal of this study: Compare the performance of these two different definitions in measuring systemic risk.

## From VaR to CoVaR

- Let  $L \sim F_L$ , continuous.  $\text{VaR}_\alpha(L) := F_L^{-1}(\alpha)$ .

**Figure 1:** Loss distribution of a univariate random variable  $L$ . Losses are given by positive numbers, gains by negative ones. The 95%-VaR-level,  $\text{VaR}_{0.95}(L)$ , is the loss, such that on average only 5% of the losses will be bigger than this. Credits McNeil et al. [2005].



## Framework and Analysis

- Stochastic framework: Bivariate model  $(X, Y) \sim F_{XY}(x, y)$ , where the random variables  $X, Y$  model the losses of two respective financial institutions.
- The main goal of CoVaR is to quantify: *What happens to  $Y$ , given that  $X$  is "in a bad state", i.e. under financial stress.*

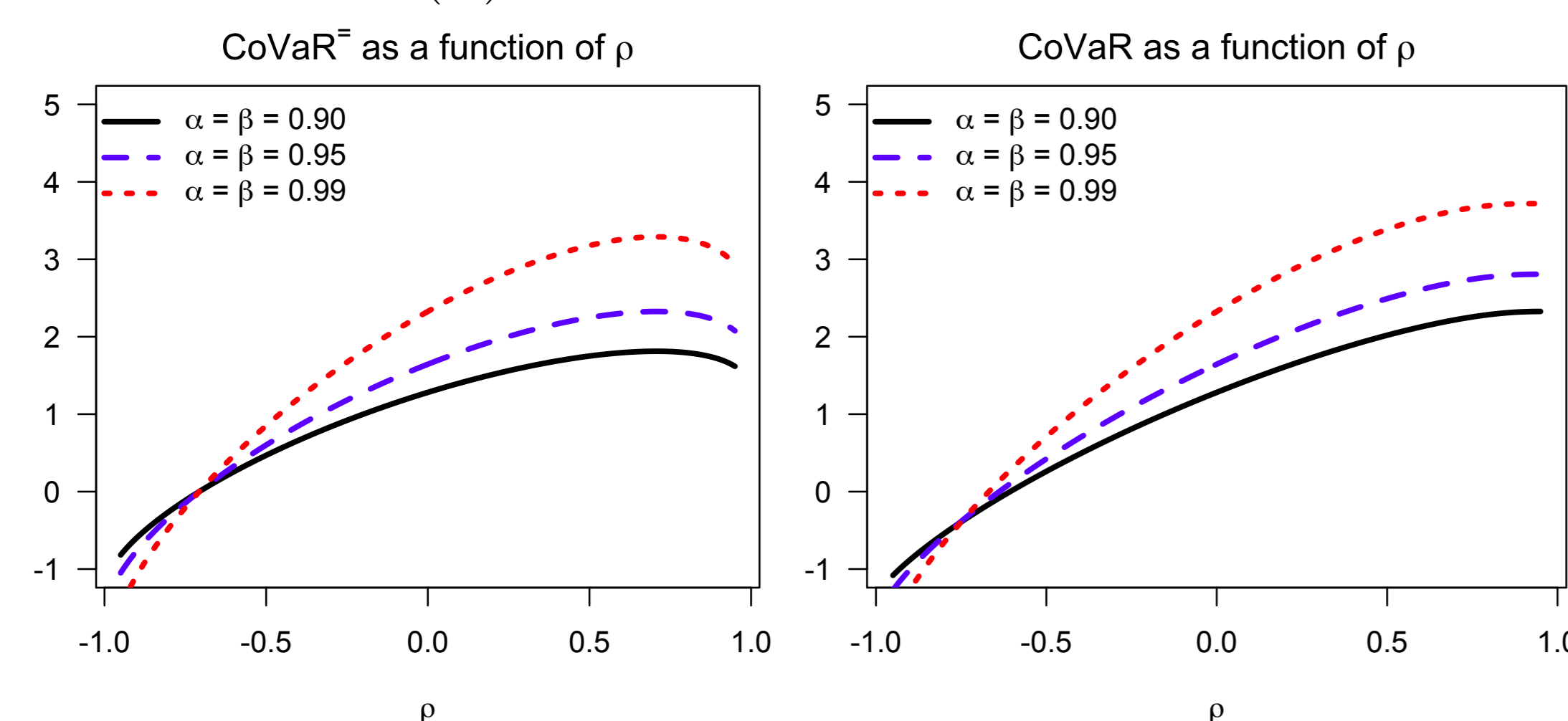
Original definition:

$$\text{CoVaR}_{\alpha, \beta}^{\text{orig}} := F_{Y|X=\text{VaR}_\alpha(X)}^{-1}(\beta).$$

Modified definition:

$$\text{CoVaR}_{\alpha, \beta} := F_{Y|X \geq \text{VaR}_\alpha(X)}^{-1}(\beta).$$

- How much of a difference does conditioning on the bad state " $X \geq \text{VaR}_\alpha(X)$ " instead of " $X = \text{VaR}_\alpha(X)$ " make?



**Figure 2:**  $\text{CoVaR}^{\text{orig}}$  seems, as opposed to  $\text{CoVaR}$ , not to be a monotonically increasing function of the dependence parameter. Capital requirements linked to this could lead to regulatory arbitrage.

## Proofs

### Non-monotonicity of $\text{CoVaR}^{\text{orig}}$

- Combining a few well known explicit formulas for the Gaussian distribution one can compute that:

$$\text{CoVaR}_{\alpha, \beta}^{\text{orig}}(Y|X) = \mu_Y + \sigma_Y \left( \rho \Phi^{-1}(\alpha) + \Phi^{-1}(\beta) \sqrt{1 - \rho^2} \right).$$

- Differentiate with respect to  $\rho$  and non-monotonicity follows immediately:

$$\partial \rho \text{CoVaR}_{\alpha, \beta}^{\text{orig}}(Y|X) = \sigma_Y \left( \Phi^{-1}(\alpha) - \frac{\rho \Phi^{-1}(\beta)}{\sqrt{1 - \rho^2}} \right).$$

### Monotonicity of $\text{CoVaR}$

A bivariate random vector  $(X, Y)$  is *elliptically distributed*  $\mathcal{E}(\mu, \Sigma, R)$  if

$$(X, Y)^T \stackrel{d}{=} \mu^T + RAW^T, \text{ where}$$

- $W = (W_1, W_2)$  is a uniformly distributed r.v. on the unit sphere  $\{x \in \mathbb{R}^2 : \|x\|_2 = 1\}$ ;
- $R$  is a non-negative random variable independent of  $W$ , called the *radial* part;
- The covariance matrix of  $(X, Y)$  is defined if and only if  $\mathbb{E}R^2 < \infty$  and can always be written as:

$$\Sigma' = \begin{pmatrix} \sigma_X^2 & \sigma_X \sigma_Y \rho' \\ \sigma_X \sigma_Y \rho' & \sigma_Y^2 \end{pmatrix}.$$

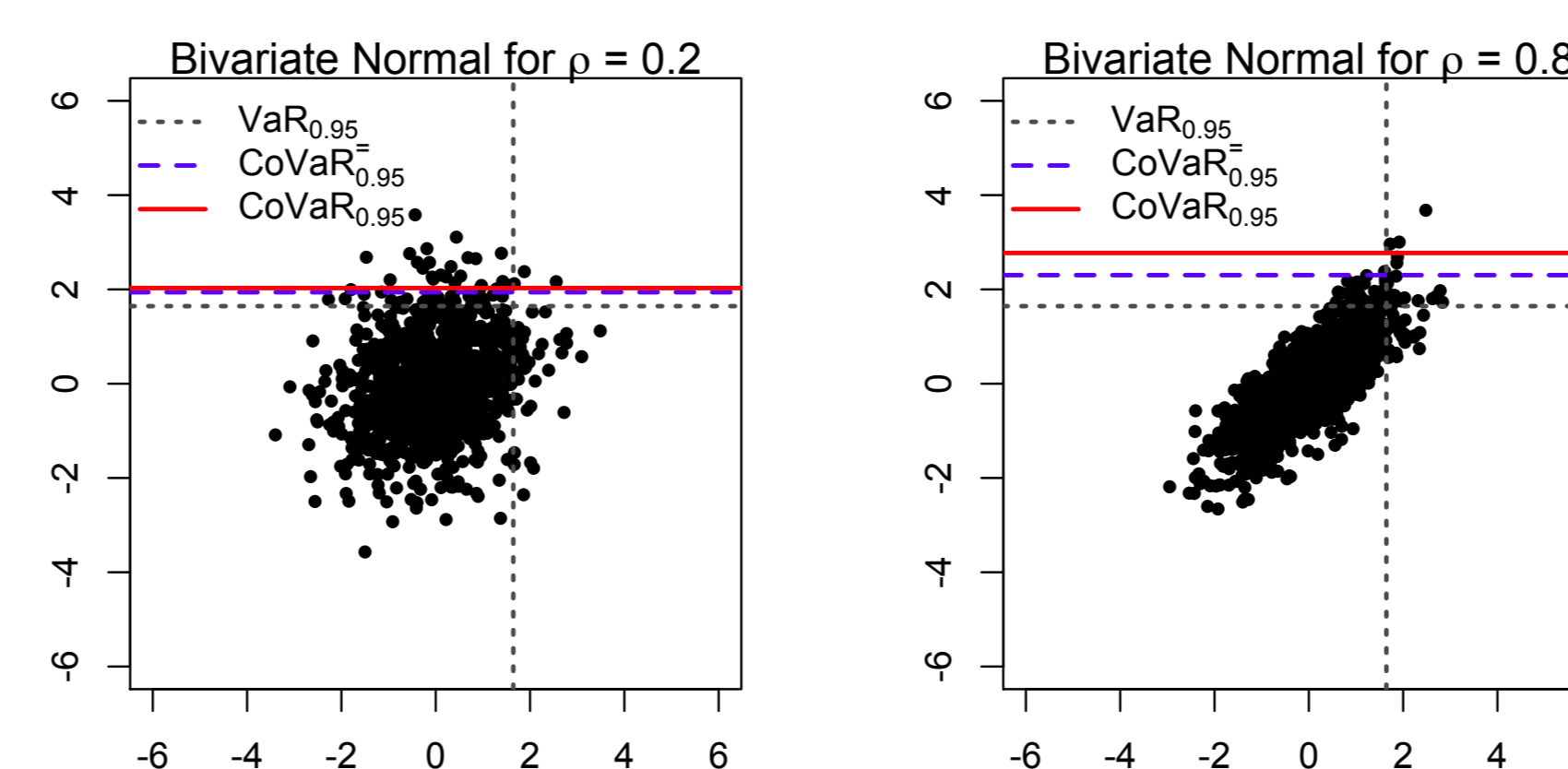
**Proposition** [M. and S., 2012]. Let  $F_{XY}$  and  $F_{X'Y'}$  have elliptical copulas with equal radial parts and dependence parameters  $\rho$  and  $\rho'$ , respectively. If  $F_X$  and  $F_{X'}$  are continuous and  $F_Y(y) \geq F_{Y'}(y)$  for all  $y \in \mathbb{R}$ , then  $\rho \leq \rho'$  implies

$$\forall \alpha, \beta \in (0, 1) \quad \text{CoVaR}_{\alpha, \beta}(Y|X) \leq \text{CoVaR}_{\alpha, \beta}(Y'|X').$$

The modified definition of  $\text{CoVaR}$  is hence consistent with the so-called *concordance ordering* of the underlying distributions and well-known results can readily be applied to a multitude of different distributions. Hence, at all confidence levels,  $\text{CoVaR}$  is an increasing function of the dependence between the components.

## Results

**Figure 3:** The probability of observing joint extremes, i.e. points falling into the upper right-hand corner, increases with stronger dependence (higher  $\rho$  here).



## Backtesting

Bound	$\rho = 0$	$\rho = 0.2$	$\rho = 0.5$	$\rho = 0.7$	$\rho = 0.9$
$\text{CoVaR}_{0.95, 0.95}^{\text{orig}}(Y X)$	0.0503	0.0601	0.0857	0.1229	0.2520
$\text{CoVaR}_{0.95, 0.95}(Y X)$	0.0503	0.0500	0.0503	0.0495	0.0499
$\text{CoVaR}_{0.99, 0.99}^{\text{orig}}(Y X)$	0.0099	0.0124	0.0189	0.0292	0.0875
$\text{CoVaR}_{0.99, 0.99}(Y X)$	0.0099	0.0101	0.0104	0.0099	0.0098

**Table 1:** Violation rates in the bivariate normal case. Monte Carlo backtesting with  $n = 10^7$  and  $\alpha, \beta \in \{0.95, 0.99\}$ .

- $\text{CoVaR}^{\text{orig}}$  fails to pick up risk when it is most pronounced and achieves a too high violation rate.
- $\text{CoVaR}^{\text{orig}}$ 's confidence level  $\beta$  can be misleading: For high-dependence scenarios, e.g.  $\rho = 0.9, \beta = 0.95$ , the level is exceeded over 25% of the instances, whereas one would expect a violation rate of  $1 - \beta \approx 5\%$ .
- By construction,  $\text{CoVaR}$  keeps a violation rate of  $1 - \beta$ .

## Conclusions

- $\text{CoVaR}^{\text{orig}}$  (and extensions of it using  $\text{CoVaR}^{\text{orig}}$  as building block) can lead to regulatory arbitrage, as they provide explicit incentives for banks to become more dependent on each other. In terms of regulatory capital one would conclude:
  - $\text{CoVaR}^{\text{orig}}$ : The more  $Y$  depends on  $X$ , the **less** capital  $Y$  requires.
  - $\text{CoVaR}$ : The more  $Y$  depends on  $X$ , the **more** capital  $Y$  requires.
- The results are even worse for non-Gaussian distributions.
- In general, dependence consistency seems to be a reasonable property to expect from systemic risk measures, and multivariate stochastic orders seem to provide a natural framework in which to analyse the behaviour of these.
- Open question: Is systemic risk measurable from market data at all? This is an important underlying assumption of the  $\text{CoVaR}$  approach.

## References

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